## COMP 677: <br> Estimation of Entropy in <br> Constant Space

Lecture 2
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## Today's lecture

- House keeping items
- Concentration of random variables
- Estimation of Entropy in Constant Space
- Feedback form


## Class project

- Projects types:
- Survey (4 papers)
- Research
- Abstract: Due 9/13 (in two weeks)
- One page
- The topic of focus
- Progress report: Due $10 / 18$
- Mid-point evaluation
- 3-page report
- Final project: Due 11/29
- 8-page final report
- Project presentation


## Next week

## Paper:

## When is Memorization of Irrelevant Training Data Necessary for High-Accuracy Learning?

Reading assignment: Due 9/6 before 4pm.

# Concentration of random variables 

## Entropy estimation in constant space

Joint work with Andrew McGregor (Umass Amherst), Jelani Nelson (UC Berkeley), Erik Waingarten (Penn)

## Estimation with memory constraints

Unknown distribution $D$
Goal: Estimate $f(D)$ with error $\epsilon$ with probability $1-\delta$ via samples

- (e.g., mean, variance, etc.)



## Estimation with memory constraints

Unknown distribution $D$
Goal: Estimate $f(D)$ with error $\epsilon$ with probability $1-\delta$ via samples

- (e.g., mean, variance, etc.)

How many samples do we need to achieve certain amount of error with limited memory?


## This work: estimating entropy

Shannon's entropy of $D=\left(p_{1}, p_{2}, \ldots, p_{n}\right)$ :

$$
H(D):=\sum_{x=1}^{n} p_{x} \log _{2} \frac{1}{p_{x}}
$$

## Entropy

Information theory :

In information theory, the entropy of a random variable is the average level of "information", "surprise", or "uncertainty" inherent to the variable's possible outcomes. Wikipedia


Entropy of a binary random variable

## This work: estimating entropy

Shannon's entropy of $D=\left(p_{1}, p_{2}, \ldots, p_{n}\right)$ :

$$
H(D):=\sum_{x=1}^{n} p_{x} \log _{2} \frac{1}{p_{x}}
$$

Used in practice to measure randomness

Applications:

- Dataset summarization
- Data compression
- Evaluating language models
- Clustering and classification


## Problem definition

Shannon's entropy of $D=\left(p_{1}, p_{2}, \ldots, p_{n}\right)$ :

$$
H(D):=\sum_{x=1}^{n} p_{x} \log _{2} \frac{1}{p_{x}}
$$



Goal:

$$
\operatorname{Pr}[|\widehat{H}-H(D)| \leq \epsilon] \geq 0.9
$$

Memory constraint: $O(1)$ words of memory (Polylog( $n, 1 / \epsilon$ ) bits)

## Our results

$n=$ domain size
$\epsilon=$ error

Theorem
There exists an algorithm for the entropy estimation problem that uses $O(1)$ words (Polylog ( $n, 1 / \epsilon$ ) bits) of memory and

$$
0\left(\frac{\eta \log (1 / \epsilon)^{4}}{\left.\epsilon_{-}^{2}\right)}\right) \text { samples. }
$$

$\Theta\left(\frac{(\hat{n})}{\epsilon \log n}+\frac{\log ^{2} n}{\left(\epsilon_{\underline{2}}^{2}\right)}\right)$ samples with no memöry constraint
[Batu, Dasgupta, Kumar, Rubinfeld. STOC 2002] [Paninski 2003] [Valiant 2008] [Valiant, Valiant. FOCS 2011] [Valiant, Valiant. JACM 2017] [Wu, Yang. IEEE Trans. IT 2016] [Jiao et al. IEEE

Trans. IT 2015] .... (and many more)
$0\left(\frac{n \log (1 / \epsilon)^{3}}{\epsilon^{3}}\right)$ samples with $O(1)$ words of memory
[Acharya, Bhadane, Indyk, Sun, NeurIPS 2019]

## A closely related model: streaming algorithms



Properties of the data stream itself

## Our results

There exists an algorithm for the entropy estimation problem that uses $O$ (1) words (Polylog ( $n, 1 / \epsilon$ ) bits) of memory and
$0\left(\frac{n \log (1 / \epsilon)^{4}}{\epsilon^{2}}\right)$ samples.

Note: Estimating the empirical entropy of the stream can NOT be done in $O(1)$ words of memory.

$$
\begin{gathered}
\Omega\left(\frac{\Gamma}{\epsilon^{2}}!_{1}^{\prime}(\log \log n+\log 1 / \epsilon)\right) \text { bits } \\
\text { [Chakrabarti, Cormode, McGregor'10] } \\
\text { [Jayaram Woodruff'19] }
\end{gathered}
$$

## Techniques

## No memory constraint

Algorithm [Valiant, Valiant'11]:

1. Compute the fingerprint of the samples

List


Number of elements


## No memory constraint

Algorithm [Valiant, Valiant'11]:

1. Compute the fingerprint of the samples
2. Come up with a histogram of a distribution that is likely to generate


Plots from [Valiant, Valiant'11]

## No memory constraint

Algorithm [Valiant, Valiant'11]:

1. Compute the fingerprint of the samples
2. Come up with a histogram of a distribution that is likely to generate
3. Output a distribution that is compatible with the histogram

Works well ignoring the labels!
Entropy
Support size

Requires memorizing all the samples

# Entropy estimation with no memory constraint 

## A simple approach

## How? Take average

$$
\begin{aligned}
& H(D):=\sum_{x=1}^{n} p_{x} \cdot \log \frac{1}{p_{x}}=\mathrm{E}_{x \sim D}\left[\log \frac{1}{p_{x}}\right] \\
& p_{x_{i}}{ }^{\prime} \text { S are unknown } \Theta x_{r} \\
& \log \frac{1}{p_{x_{1}}} \log \frac{1}{p_{x_{2}}} \quad \ldots \quad \log \frac{1}{p_{x_{r}}}
\end{aligned}
$$

## Estimate probabilities

Fix $m$. Count $i$ 's in next $m$ samples.

$$
\text { Set } \hat{p}_{x}=\frac{\# \text { instances }}{m}
$$

$$
\text { \#instances of } x \sim \operatorname{Bin}\left(m, p_{x}\right)
$$



In the example: $\frac{2}{6}$

## How? Take average

$$
\begin{align*}
& H(D):=\sum_{x=1}^{n} p_{x} \cdot \log \frac{1}{p_{x}}=\mathrm{E}_{x \sim D}\left[\log \frac{1}{\hat{p}_{x}}\right] \\
& \log \frac{1}{\hat{p}_{\hat{x}_{1}}} \log \frac{1}{\hat{p}_{\hat{x}_{2}}} \\
& \frac{1}{r} \sum_{i=1}^{r} \log \frac{1}{\hat{p}_{\hat{x}_{1}}} \xrightarrow{\text { large } \mathrm{r}}
\end{align*}
$$

## How? Take average

$$
\begin{aligned}
& \frac{1}{r} \sum_{i=1}^{r} \log \frac{1}{\frac{1}{\hat{p}_{x_{1}}}} \xrightarrow{\text { large } \mathrm{r}} \mathrm{E}_{x \sim D}\left[\log \frac{1}{\left.\frac{1}{\hat{p}_{x_{1}}}\right]} \xrightarrow{\text { large } \mathrm{m}} \mathrm{E}_{x \sim D}\left[\log \frac{1}{p_{x}}\right]=\mathrm{H}(\mathrm{D})\right. \\
& \text { Error of estimation Bias } \\
& E\left[\# \text { samples] }=\Theta(r \cdot m)=\Theta\left(\frac{n \log \left(\frac{n}{\epsilon}\right)}{\epsilon^{3}}\right)\right.
\end{aligned}
$$

## Analysis of error

$$
\begin{aligned}
& n=\text { domain size } \\
& \epsilon=\text { error } \\
& m=\text { number of samples to estimate } p_{i} \\
& r=\text { number of rounds }
\end{aligned}
$$

Error: $|H(D)-\widehat{H}| \leq \epsilon$

$$
|H(D)-\widehat{H}| \leq\left|H(D)-\mathrm{E}\left[\widehat{H}_{i}\right]\right|+\left|\mathrm{E}\left[\widehat{H}_{i}\right]-\widehat{H}\right|
$$

$$
\leq \underbrace{\left|\mathrm{E}_{i \sim D}\left[\log \frac{1}{p_{i}}\right]-\mathrm{E}_{i \sim D}\left[\log \frac{1}{\hat{p}_{i}}\right]\right|}_{\text {Bias }}+\underbrace{\left|\mathrm{E}\left[\widehat{H}_{i}\right]-\widehat{H}\right|}_{\text {Error of estimation }}
$$

$m>\Omega(n / \epsilon)$ implies bias $<\epsilon / 2 r=\Theta\left(\log m / \epsilon^{2}\right)$ implies that error $<\epsilon / 2$

$$
E[\# \text { samples }]=\Theta(r \cdot m)=\Theta\left(\frac{n \log \left(\frac{n}{\epsilon}\right)}{\epsilon^{3}}\right)
$$

Simple algorithm [Plug-in estimator]

$$
H(D):=\sum_{i=1}^{n} p_{i} \cdot \log 1 / p_{i}=\mathrm{E}_{i \sim D}\left[\log 1 / p_{i}\right]
$$

1. Repeat $r$ times
2. Draw $i \sim D$.

Fix $m$. Count $i$ 's in next $m$ samples.
2. $\hat{p}_{i} \leftarrow$ Estimate $p_{i}$
3. $\widehat{H}_{i} \leftarrow \log 1 / \hat{p}_{i}$

$$
\# \text { instances of } i \sim \operatorname{Bin}\left(m, p_{i}\right)
$$

Set $\hat{p}_{i}=\frac{\# \text { instances }}{m}$
In the example: $\frac{2}{6}$


## Simple algorithm

$$
\begin{aligned}
& n=\text { domain size } \\
& \epsilon=\text { error } \\
& m=\text { number of samples to estimate } p_{i} \\
& r=\text { number of rounds }
\end{aligned}
$$

$$
H(D):=\sum_{i=1}^{n} p_{i} \cdot \log 1 / p_{i}=\mathrm{E}_{i \sim D}\left[\log 1 / p_{i}\right]
$$

1. Repeat $r$ times
2. Draw $i \sim D$.
3. $\hat{p}_{i} \leftarrow$ Estimate $p_{i}$
4. $\widehat{H}_{i} \leftarrow \log 1 / \hat{p}_{i}$
5. Output: $\widehat{H}:=\frac{1}{\mathrm{r}} \sum_{i=1}^{r} \widehat{H}_{i}$

Fix $m$. Count the number of instances of $i$ in the next $m$ samples.


## Idea I: Estimate via negative binomials

Count the number of samples until $t$ instances of $x$ are observed.

$$
\begin{gathered}
\text { \#samples } \sim \text { Negative } \operatorname{Bin}\left(t, p_{x}\right) \\
\text { Set } X_{x}=\frac{\# \text { samples }}{t} \\
\mathrm{E}\left[X_{x}\right]=1 / p_{x}
\end{gathered}
$$

In the example for $t=2: X_{x}=\frac{7}{2}$


Analysis of error
?
$n=$ domain size of the distribution
$\epsilon=$ error parameter
$r=$ number of rounds
$t=$ number of observed instance of $i$

Error: $|H(D)-\widehat{H}| \leq \epsilon$
$|H(D)-\widehat{H}| \leq\left|H(D)-\mathrm{E}\left[\widehat{H}_{i}\right]\right|+\left|\mathrm{E}\left[\widehat{H}_{i}\right]-\widehat{H}\right|$

$$
\leq \underbrace{\left\lvert\, \mathrm{E}_{i \sim D}\left[\log \frac{1}{p_{i}}\right]-\mathrm{E}_{i \sim D}\left[\log \frac{1}{\hat{p}_{i}}\right]\right.}_{\text {Bias }}+\underbrace{\left|\mathrm{E}\left[\hat{H}_{i}\right]-\widehat{H}\right|}_{\text {Error of }}
$$

$t=\Theta(1 / \epsilon)$ implies bias $<\epsilon / 2 \quad r=\Theta\left(\log ^{2} n / \epsilon^{2}\right)$ implies that error $<\epsilon / 2$

$$
E[\# \text { samples }]=\Theta(r \cdot t \cdot n)=\Theta\left(n \log ^{2} n / \epsilon^{3}\right)
$$

$n=$ domain size of the distribution
$\epsilon=$ error parameter
$r=$ number of rounds
$t=$ number of observed instance of $i$
$X_{i}=$ number of samples to see $t$ instance of $i$
Idea: Estimate bias and subtract it from $\widehat{H}$.

Let $Y_{i} \leftarrow p_{i} X_{i}$
Bias $=\left|\mathrm{E}_{i \sim D}\left[\log 1 / p_{i}\right]-\mathrm{E}_{i \sim D}\left[\log X_{i}\right]\right|=\left|\mathrm{E}_{i \sim D}\left[\log Y_{i}\right]\right|$
$\mathrm{E}_{i \sim D}\left[Y_{i}\right]=1$. Taylor expansion around $\mathrm{Y}=1$ :
Bias $=\mathrm{E}_{i \sim D}\left[\log Y_{i}\right]=\mathrm{E}\left[Y_{i}-1-\frac{\left(Y_{i}-1\right)^{2}}{2}+\frac{\left(Y_{i}-1\right)^{3}}{3}-\cdots\right]$

## Idea II: Remove bias

Idea: Truncated Taylor expansion. Keep the first $s=\log (1 / \epsilon)$ terms.

$\operatorname{Pr}[\mathrm{k}$ samples are equal $]=p_{i}^{k}$

```
                                    n = domain size of the distribution
                                    \epsilon= error parameter
                                    r = number of rounds
                            t= number of observed instance of i
                            Xi}=\mathrm{ number of samples to see t
                        instance of }
\(\mathrm{E}\left[X_{i}\right]=1 / p_{i}\)
```

Idea: Bucketing
Partition the range of $X_{i}$ into $L$ intervals

$$
\mathrm{E}_{i \sim D}\left[\log X_{i}\right]=\sum_{\ell=1}^{L} \underbrace{\operatorname{Pr}\left[X_{i} \in I_{\ell}\right]} \underbrace{\mathrm{E}\left[\log X_{i} \mid X_{i} \in I_{\ell}\right]}_{1}
$$

Estimate $\widehat{q}_{L}$ and $\widehat{H}_{L}$
$q_{\ell} \quad H_{\ell}$


## $n=$ domain size of the distribution <br> $\epsilon=$ error parameter <br> Idea III: Remove $\log n$ factors <br> $r=$ number of rounds <br> $t=$ number of observed instance of $i$ <br> $X_{i}=$ number of samples to see $t$ instance of $i$ <br> $\mathrm{E}\left[X_{i}\right]=1 / p_{i}$

Error $\leq\left|\sum_{\ell=1}^{L-1}\left(\hat{q}_{\ell}-q_{\ell}\right) \cdot\left(H_{\ell}-H_{L}\right)\right|+\left|\sum_{\ell=1}^{L} q_{\ell} \cdot\left(H_{\ell}-\widehat{H}_{\ell}\right)\right|$
Bucke Removing $O(\log n)$. zuracy.

$$
\text { Removing } O(\log n) \text {. }
$$

$L$ Buckets

