# COMP 677: Estimation of Entropy in

## **Constant Space**

Lecture 2

Maryam Aliakbarpour

Fall 2023

#### Today's lecture

- House keeping items
- Concentration of random variables
- Estimation of Entropy in Constant Space
- Feedback form

#### Class project

- Projects types:
  - Survey (4 papers)
  - Research
- Abstract: Due 9/13 (in two weeks)
  - One page
  - The topic of focus
- Progress report: Due 10/18
  - Mid-point evaluation
  - 3-page report
- Final project: Due 11/29
  - 8-page final report
- Project presentation

#### Next week

Paper:

When is Memorization of Irrelevant Training Data Necessary for High-Accuracy Learning?

Reading assignment: Due 9/6 before 4pm.

# Concentration of random variables

# Entropy estimation in constant space

Joint work with Andrew McGregor (Umass Amherst), Jelani Nelson (UC Berkeley), Erik Waingarten (Penn)

### Estimation with memory constraints

Unknown distribution D

Goal: Estimate f(D) with error  $\epsilon$  with probability  $1 - \delta$  via samples

- (e.g., mean, variance, etc.)



### Estimation with memory constraints

#### Unknown distribution D

Goal: Estimate f(D) with error  $\epsilon$  with probability  $1 - \delta$  via samples

- (e.g., mean, variance, etc.)

How many samples do we need to achieve certain amount of error with limited memory?





#### This work: estimating entropy

Shannon's entropy of  $D = (p_1, p_2, \dots, p_n)$ :

Entropy



Feedback

Entropy of a binary random variable

### This work: estimating entropy

Shannon's entropy of  $D = (p_1, p_2, ..., p_n)$ :

$$H(D) \coloneqq \sum_{x=1}^{n} p_x \log_2 \frac{1}{p_x}$$

Used in practice to measure randomness

Applications:

- Dataset summarization
- Data compression
- Evaluating language models
- Clustering and classification

|--|--|

#### Problem definition

Shannon's entropy of  $D = (p_1, p_2, ..., p_n)$ :



Goal:

$$\Pr[\left|\widehat{H} - H(D)\right| \le \epsilon] \ge 0.9$$

Memory constraint: O(1) words of memory  $(Polylog(n, 1/\epsilon)$  bits)

n =domain size  $\epsilon =$ error

#### Our results

[A, McGregor, Nelson, Waingarten'22] Theorem There exists an algorithm for the entropy estimation problem that uses O(1)words  $(Polylog(n, 1/\epsilon)$  bits) of memory and  $0\left(\frac{n\log(1/\epsilon)^4}{\epsilon^2}\right)$  samples.  $\Theta\left(\frac{n}{\epsilon \log n} + \frac{\log^2 n}{\epsilon^2}\right)$  samples with no  $0\left(\frac{n\log(1/\epsilon)^3}{\epsilon^3}\right)$  samples with O(1) words of memory memory constraint [Acharya, Bhadane, Indyk, Sun, NeurIPS] [Batu, Dasgupta, Kumar, Rubinfeld. STOC 2002] [Paninski 2003] [Valiant 2008] [Valiant, Valiant. FOCS 2011] [Valiant, Valiant. 2019] JACM 2017] [Wu, Yang. IEEE Trans. IT 2016] [Jiao et al. IEEE Trans. IT 2015] .... (and many more)

#### A closely related model: streaming algorithms



n = domain size  $\epsilon =$  error

#### Our results

Theorem [A, McGregor, Nelson, Waingarten'22] There exists an algorithm for the entropy estimation problem that uses O(1) words  $(Polylog(n, 1/\epsilon)$  bits) of memory and  $O\left(\frac{n \log(1/\epsilon)^4}{\epsilon^2}\right)$  samples.

Note: Estimating the empirical entropy of the stream can NOT be done in O(1) words of memory.

$$\Omega\left(\frac{1}{\epsilon^2}, (\log \log n + \log 1/\epsilon)\right)$$
 bits

lakrabarti, Cormode, McGregor 10 [Jayaram Woodruff'19]

# Techniques

#### No memory constraint

Algorithm [Valiant, Valiant'11]:

1. Compute the fingerprint of the samples







#### No memory constraint

Algorithm [Valiant, Valiant'11]:

- 1. Compute the fingerprint of the samples
- 2. Come up with a histogram of a distribution that is likely to generate



Plots from [Valiant, Valiant'11]

#### No memory constraint

Algorithm [Valiant, Valiant'11]:

- 1. Compute the fingerprint of the samples
- 2. Come up with a histogram of a distribution that is likely to generate
- 3. Output a distribution that is compatible with the histogram

Works well ignoring the labels! Entropy

Support size

Requires memorizing all the samples





Entropy estimation with no memory constraint

# A simple approach

 $\epsilon = error$ How? Take average  $H(D) \coloneqq \sum p_x \cdot \log \frac{1}{p_x} = \mathbb{E}_{x \sim D} \left[ \log \frac{1}{p_x} \right]$  $p_{x_i}$ 's are unknown  $\Theta$  $x_r$  $x_1$  $\log \frac{1}{p_{x_1}} \log \frac{1}{p_{x_2}}$  $\log \frac{1}{p_{x_r}}$ ...  $\sum_{i=1}^{r} \log \frac{1}{p_{x_i}} \xrightarrow{\text{large } r} E_{x \sim D} \left[ \log \frac{1}{p_x} \right]$ = H(D)

n =domain size

#### Estimate probabilities

Fix m. Count i's in next m samples.

Set  $\hat{p}_x = \frac{\# \text{ instances}}{m}$ 

#instances of  $x \sim Bin(m, p_x)$ 



In the example:  $\frac{2}{6}$ 



n =domain size  $\epsilon =$ error

#### How? Take average



#### n =domain size $\epsilon = error$ Analysis of error m = number of samples to estimate $p_i$ r = number of rounds $\begin{array}{l}?\\ \text{Error:} \left| H(D) - \widehat{H} \right| \leq \epsilon \end{array}$ $|H(D) - \widehat{H}| \le |H(D) - \mathbb{E}[\widehat{H}_i]| + |\mathbb{E}[\widehat{H}_i] - \widehat{H}|$ $\leq \left| \mathbf{E}_{i \sim D} \left[ \log \frac{1}{p_i} \right] - \mathbf{E}_{i \sim D} \left[ \log \frac{1}{\hat{p}_i} \right] \right| + \left| \mathbf{E} \left[ \hat{H}_i \right] - \hat{H} \right|$ Bias Error of estimation $m > \Omega(n/\epsilon)$ implies bias $< \epsilon/2$ $r = \Theta(\log m/\epsilon^2)$ implies that error $< \epsilon/2$ $E[\text{#samples}] = \Theta(r \cdot m) = \Theta\left(\frac{n\log\left(\frac{n}{\epsilon}\right)}{\epsilon^3}\right)$

n =domain size  $\epsilon =$ error

## Simple algorithm [Plug-in estimator]

$$H(D) \coloneqq \sum_{i=1}^{n} p_i \cdot \log 1/p_i = \mathrm{E}_{i \sim D}[\log 1/p_i]$$

- 1. Repeat r times
  - 1. Draw  $i \sim D$ .
  - 2.  $\hat{p}_i \leftarrow \text{Estimate } p_i \dots$
  - 3.  $\hat{H}_i \leftarrow \log 1/\hat{p}_i$
- 2. Output:  $\widehat{H} := \frac{1}{r} \sum_{i=1}^{r} \widehat{H}_i$

Fix m. Count i's in next m samples.

#instances of 
$$i \sim Bin(m, p_i)$$
  
Set  $\hat{p}_i = \frac{\# instances}{m}$   
In the example:  $\frac{2}{6}$ 

#### Simple algorithm

$$H(D) \coloneqq \sum_{i=1}^{n} p_i \cdot \log 1/p_i = \mathbb{E}_{i \sim D}[\log 1/p_i]$$

- 1. Repeat *r* times
  - 1. Draw  $i \sim D$ .
  - 2.  $\hat{p}_i \leftarrow \text{Estimate } p_i \dots$
  - 3.  $\hat{H}_i \leftarrow \log 1/\hat{p}_i$
- 2. Output:  $\widehat{H} := \frac{1}{r} \sum_{i=1}^{r} \widehat{H}_i$

Fix m. Count the number of instances of i in the next m samples.

n =domain size

r = number of rounds

m = number of samples to estimate  $p_i$ 

 $\epsilon = error$ 

#instances of 
$$i \sim Bin(m, p_i)$$
  
Set  $\hat{p}_i = \frac{\# instances}{m}$   
In the example:  $\frac{2}{6}$ 

#### Idea l: Estimate via negative binomials

Count the number of samples until t instances of x are observed.

#samples ~ Negative Bin  $(t, p_x)$ Set  $X_x = \frac{\text{# samples}}{t}$  $E[X_x] = 1/p_x$ 

In the example for 
$$t = 2$$
:  $X_x = \frac{7}{2}$ 



#### n = domain size of the distribution $\epsilon = \text{error parameter}$ Analysis of error r = number of rounds t = number of observed instance of *i* $\begin{array}{l} ?\\ \mathsf{Error:} \left| H(D) - \widehat{H} \right| \leq \epsilon \end{array}$ $|H(D) - \widehat{H}| \le |H(D) - \mathbb{E}[\widehat{H}_i]| + |\mathbb{E}[\widehat{H}_i] - \widehat{H}|$ $\leq \left| \mathbf{E}_{i \sim D} \left[ \log \frac{1}{p_i} \right] - \mathbf{E}_{i \sim D} \left[ \log \frac{1}{\hat{p}_i} \right] \right| + \left| \mathbf{E} \left[ \hat{H}_i \right] - \hat{H} \right|$ Bias Error of estimation

 $t = \Theta(1/\epsilon) \text{ implies bias} < \epsilon/2 \quad r = \Theta(\log^2 n/\epsilon^2) \text{ implies that error} < \epsilon/2$  $E[\text{#samples}] = \Theta(r \cdot t \cdot n) = \Theta(n \log^2 n/\epsilon^3)$ 

#### Idea II: Remove bias

Idea: Estimate bias and subtract it from  $\widehat{H}$ .

Let  $Y_i \leftarrow p_i X_i$ Bias =  $|E_{i \sim D}[\log 1/p_i] - E_{i \sim D}[\log X_i]| = |E_{i \sim D}[\log Y_i]|$ 

 $E_{i\sim D}[Y_i] = 1$ . Taylor expansion around Y = 1: Bias =  $E_{i\sim D}[\log Y_i] = E\left[Y_i - 1 - \frac{(Y_i - 1)^2}{2} + \frac{(Y_i - 1)^3}{3} - \cdots\right]$ 

n = domain size of the distribution  $\epsilon = \text{error parameter}$  r = number of rounds t = number of observed instance of i  $X_i = \text{number of samples to see } t$ instance of i $E[X_i] = 1/p_i$ 

#### Idea II: Remove bias

Idea: Truncated Taylor expansion. Keep the first  $s = \log(1/\epsilon)$  terms.



Idea III: Remove log n factors

Idea: Bucketing Partition the range of  $X_i$  into L intervals

Estimate  $\hat{q}_L$  and  $\hat{H}_L$ 

n = domain size of the distribution  $\epsilon = \text{error parameter}$  r = number of rounds t = number of observed instance of i  $X_i = \text{number of samples to see } t$ instance of i $E[X_i] = 1/p_i$ 

H₽



 $\mathbb{E}_{i \sim D} \left[ \log X_i \right] = \sum_{\ell=1}^{L} \Pr[X_i \in I_\ell] : \mathbb{E}[\log X_i | X_i \in I_\ell]$ 

 $q_\ell$ 

## Idea III: Remove log n factors

n = domain size of the distribution  $\epsilon = \text{error parameter}$  r = number of rounds t = number of observed instance of i  $X_i = \text{number of samples to see } t$ instance of i $E[X_i] = 1/p_i$ 

$$\operatorname{Error} \leq \left| \sum_{\ell=1}^{L-1} (\hat{q}_{\ell} - q_{\ell}) \cdot (H_{\ell} - H_{L}) \right| + \left| \sum_{\ell=1}^{L} q_{\ell} \cdot \left( H_{\ell} - \widehat{H}_{\ell} \right) \right|$$

